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Education and Health: Redrawing the Preston Curve

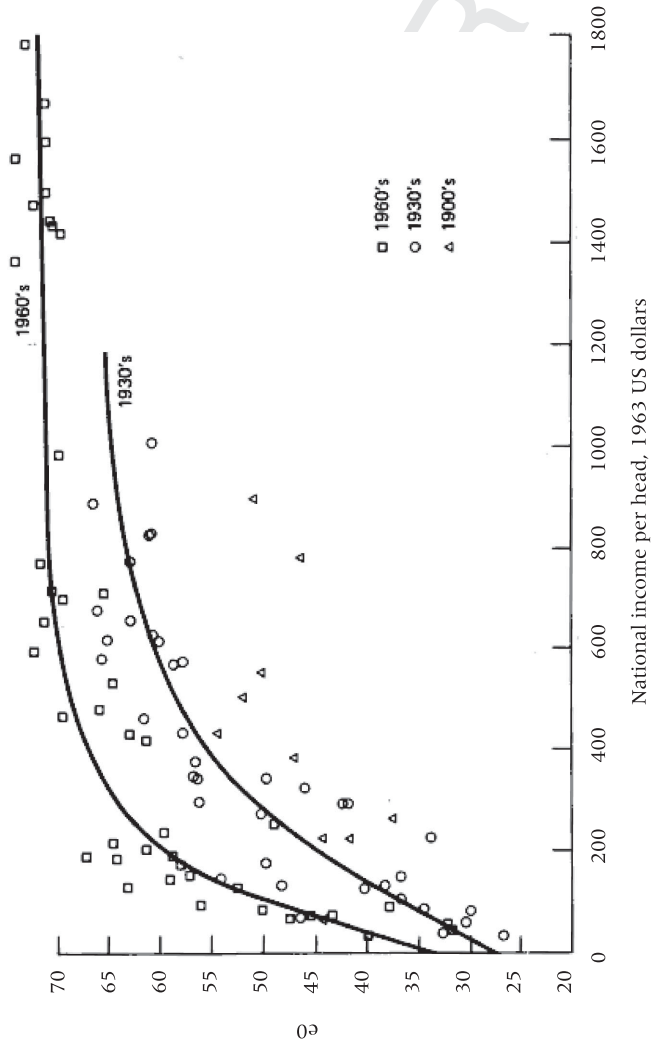
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PROGRESS IN HUMAN health and life expectancy is closely associated with socioeconomic development. Better nutrition and greater affordability of health care associated with higher income have been widely considered as primary determinants of historical and contemporary mortality declines. McKeown's (1976) influential book on the modern rise of population attributed the secular mortality decline largely to improving standards of living. Reviewing mortality improvements in Britain during the second half of the nineteenth century and the beginning of the twentieth, he argued that medical discoveries were of little consequence for the significant gains in survival during this period. His analysis served as a reference point of Preston's (1975) article, which is the focus of the present study. Preston showed that the global pattern over the twentieth century indicates an upward shift of the curve that links GDP per person on the horizontal axis and life expectancy on the vertical (Figure 1).¹ Preston interpreted this shift as the effect of medical progress and health care over and above the effect of income. In many of the studies of this issue that followed Preston's lead, the assumption that income is the most important driver of mortality decline has been an unquestioned starting point.

A very different picture was drawn by Caldwell in a 1986 article on routes to low mortality in poor countries. Based on a major Rockefeller Foundation study on Kerala (India), Sri Lanka, and Costa Rica, Caldwell discussed the factors that led to breakthrough mortality declines in those populations as opposed to others and identified "female autonomy," which he saw largely as a function of female education, as the single most important factor, together with efficient local health services. He also stated that his conclusion that low mortality does not come as an unplanned spinoff from economic growth was "out of step with today's dominant economic and political ideologies in the development field" (Caldwell 1986, p. 209). And this still seems to be the case three decades later, despite the fact that more recent research points to the overriding importance of education and the associated cognitive changes affecting risk perception, planning

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FIGURE 1 Scatter-diagram of relations between life expectancy at birth (e0) and national income per head for nations in the 1900s, 1930s, and 1960s



2 horizon, and access to information promoting health-related behaviors and
3 use of health care facilities (Baker et al. 2011; Lutz and Skirbekk 2014).

4 The question whether income or education is the more important de-
5 terminant of global health and mortality decline is relevant for setting pol-
6 icy priorities in both developing and industrialized countries. The answer is
7 of immediate concern in choosing between programs that directly promote
8 economic growth and those that focus on enhancing school enrollment and
9 quality of schooling. In an ideal world one would choose both in addition
10 to good local health services but, in reality, even in rich countries there are
11 budgetary constraints that require policymakers to set priorities. For help-
12 ing to set these priorities, it is necessary to assess the relative importance of
13 both factors. The question of relative importance is the focus of what fol-
14 lows. We address the issue at the macro level by plotting a modification of
15 the Preston curve in which GDP per person is replaced by mean years of
16 schooling among the adult population. This is done both for life expectancy
17 at birth and for child mortality, and in both cases educational attainment ex-
18 plains the pattern better than GDP per person. This redrawing of the curve
19 is complemented by a multivariate analysis to quantitatively assess the
20 relative difference of the two effects.
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23 The Preston curve and its perception 24

25 In 1975 Samuel Preston published an influential paper, “The changing re-
26 lation between mortality and level of economic development,” in which he
27 plots the global relationship between GDP per capita and life expectancy at
28 different points in time. He finds that over time the curve that depicts the
29 relationship has moved upward, implying that a similar level of income is
30 associated with higher life expectancy at later points in time (Preston 1975).
31 He attributes this extra gain to medical progress. This view of the relation-
32 ship, known as the Preston curve, has since become widely cited. Although
33 Preston was cautious in interpreting this relationship as an association and
34 not necessarily a causal one, subsequent interpretations of this relation of-
35 ten had no doubt that it was based on direct causation (e.g., Pritchett and
36 Summers 1996).

37 In a more elaborate follow-up study, Preston (1980) introduced liter-
38 acy and calorie supply in addition to GDP per person into regression models
39 to explain differences in levels of life expectancy for 36 countries in 1940
40 and 120 countries in 1970. The estimated coefficients showed similar pat-
41 terns for both times, and literacy was highly significant. In Preston’s words:
42 “The coefficients indicate that a 10 percentage point increase in literacy is
43 associated at both points with a gain in life expectancy of approximately 2
44 years, and that a 10% gain in national income by itself increases life ex-
45 pectancy by approximately one-half year” (Preston 1980, p. 306). These

2 interesting findings, however, were largely overlooked in subsequent re-
3 search on development and mortality.

4 This apparent positive association between income and health has
5 given rise to the widespread view among development economists that in-
6 creased wealth leads causally to increased health. In the extreme, Pritchett
7 and Summers (1996) argue that focusing on economic growth in develop-
8 ing countries will lead directly to reductions in infant mortality rates and
9 gains in life expectancy. While many other economists hold a less defini-
10 tive position and acknowledge drivers in addition to income, the possibility
11 that the apparent empirical association between income and health could
12 be largely spurious and not of a causal nature has not been considered in
13 this body of economic literature.

14 In 2007 an issue of the *International Journal of Epidemiology* was devoted
15 to a reprint of Preston's 1975 article and several comments by distinguished
16 scholars in the field (IEA 2007). In the contribution most relevant for our
17 analysis, Bloom and Canning (2007) revisit the Preston curve and state that
18 his paper "remains a cornerstone of both global public health policy and
19 academic discussion of public health. Preston's paper illuminates two cen-
20 tral 'stylized facts'. The first is a strong, positive relationship between na-
21 tional income levels and life expectancy in poorer countries, though the
22 relationship is non-linear as life expectancy levels in richer countries are
23 less sensitive to variations in average income. The second is that the re-
24 lationship is changing, with life expectancy increasing over time at all in-
25 come levels. ... Although the basic facts set out by Preston are generally
26 accepted, there is still a great deal of dispute about the mechanisms that
27 lie behind the relationships and the policy implications we can draw from
28 them" (p. 498).

29 Bloom and Canning also point to another body of literature in pub-
30 lic health in which, for example, Cutler, Deaton, and Lleras-Muney (2006)
31 conclude that scientific and technical advances should be seen as "the ul-
32 timate determinant of health." A further argument against focusing on in-
33 come growth as the primary method of reducing the burden of ill health
34 lies in the apparently very weak temporal association between periods of
35 economic growth and periods of improvement in population health, sug-
36 gesting that if the relationship were causal, it has long and variable lags.
37 While rising incomes imply greater resources for society, these resources
38 need not necessarily be spent in ways that improve health. While Bloom
39 and Canning do not question the basic assumption that income growth and
40 health are closely linked, they add a cautionary sentence that is the starting
41 point for our study: "Although there is a strong case for the direct effect of
42 income on health due to nutrition and health interventions becoming more
43 affordable, it may be that income is also acting as a proxy for a wider mea-
44 sure of socioeconomic status and development and that the causal effect is
45 due to other mechanisms, for example, education" (p. 498)

2 In a more recent assessment of the Preston curve, Mackenbach and
3 Looman (2013) find that for European countries increases in life expectancy
4 after 1960 have been accompanied by a much smaller upward shift in the
5 curve than previously. They attribute this to a changing pattern of causes
6 of death away from infectious diseases and conclude that “declines in mor-
7 tality from cardiovascular disease were mainly attributable to increases in
8 national income.” This conclusion seems to be reached simply by eliminat-
9 ing medical progress as a dominating reason and assuming that the only
10 other possible determinant of mortality was income, which thus should be
11 the cause of this decline. The possibility of another driver that jointly deter-
12 mines income and health and causes their correlation was not considered.

13 In this article, we address the hypothesis that the apparent statisti-
14 cal association between income and health—as described by the Preston
15 curve—could in fact be a largely spurious association resulting from the fact
16 that improving educational attainment is a key determinant of both better
17 health and rising incomes. Before doing so, we discuss the issue of causality.
18

20 On causality

21
22 The question of the causal nature of the effects of education and income
23 on health has attracted much attention and controversy. It is a question
24 that needs to be addressed in order to rule out the possibility that the as-
25 sociations that are being interpreted could also be spurious. The question
26 has important policy implications. If, for instance, the empirical association
27 between income and health is not directly causal but rather due to a third
28 factor such as education, then an increase in income—e.g., through poli-
29 cies directly aiming at economic growth—would not result in the expected
30 health improvements unless educational attainment also improved simul-
31 taneously. The same is true *mutatis mutandis* for the association between
32 educational attainment and health.

33 Causality in the social sciences has to be viewed differently than in
34 the natural sciences because human behavior is culturally embedded and
35 what is found to be direct causation of behavior in a given setting cannot be
36 assumed to have universal predictive power for all societies and all times.
37 Inspired by the comprehensive review of causality in demography by Ní
38 Bhrolcháin and Dyson (2007), Lutz and Skirbekk (2014) introduced the no-
39 tion of “functional causality” in the context of “intervention sciences” (Lutz
40 and Striessnig 2015; Pearl 2000). Intervention sciences are the social and
41 economic sciences that try to understand how the most important forces
42 of change in a society function in order to predict the future evolution of
43 the social system. Such conditional predictions about future trends can be
44 based on the assumption of no intervention or of alternative interventions
45 and their likely consequences.

2 To establish functional causality, three criteria have to be met: (i)
3 there must be strong empirically observed associations between the two
4 factors studied; (ii) there must be a plausible narrative about the mecha-
5 nisms through which one force influences the other; and (iii) other obvious
6 competing explanations of the observed association should be ruled out.
7 Examples of such competing explanations are self-selection, reverse causal-
8 ity, and joint determination by a third force (Lutz and KC 2011; Lutz and
9 Skirbekk 2014). Lutz and Skirbekk (2014) give a comprehensive overview
10 of dozens of relevant studies on the topic and conclude that it is justified to
11 assume functional causality for the global-level relationship between educa-
12 tional attainment and health/mortality over the twentieth and twenty-first
13 centuries.

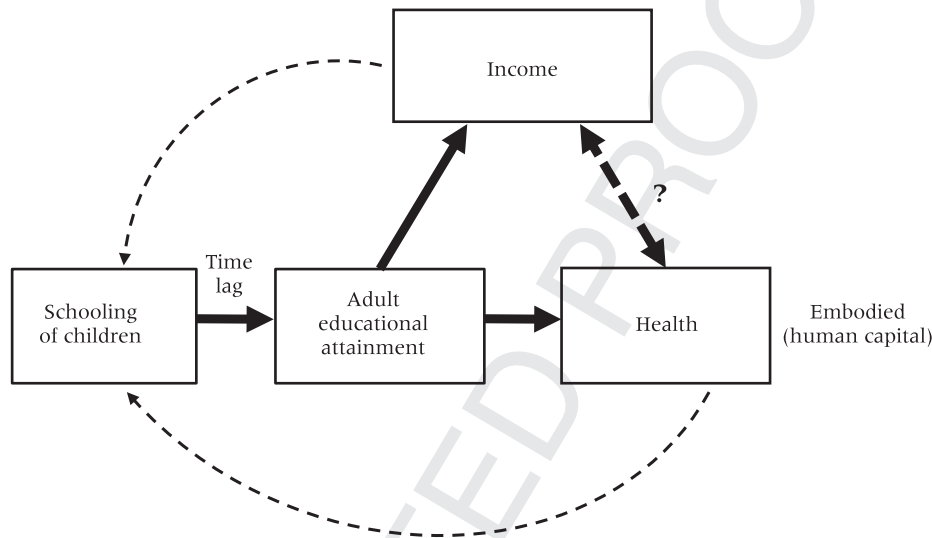
14 Several global assessments of the relationship between health and
15 mortality (Baker et al. 2011; KC and Lentzner 2010; Pamuk, Fuchs, and
16 Lutz 2011) show that, on all continents and at different levels of socio-
17 economic development, the less-educated segments of the population have
18 significantly higher mortality and morbidity than those who are better ed-
19 ucated. In virtually all countries, children of better-educated mothers ex-
20 perience lower mortality. But since better-educated people also tend to
21 live in richer households, the question arises: what is more important for
22 child survival in developing countries, mothers' education or household in-
23 come/wealth? This question has been the focus of studies using the largest
24 available individual-level data set by pooling the samples of Demographic
25 and Health Surveys (DHS) in 43 developing countries (Fuchs, Pamuk, and
26 Lutz 2010; Pamuk, Fuchs, and Lutz 2011). Using multi-level regression
27 models, analysis of the relative effects of mother's education and economic
28 resources on infant mortality at the family, community, and country levels
29 shows that the effect of education clearly dominates over income/wealth
30 at all levels. The empirical evidence is equally compelling for the effect of
31 education on adult mortality. In virtually all countries for which data exist,
32 better-educated people have higher life expectancies (Caselli et al. 2014).
33 The differences vary in extent among countries and are generally greater
34 for men than for women. Among industrialized countries, differences tend
35 to be lowest in Southern European countries and are highest in Eastern
36 Europe. In Russia, the difference between the highest and lowest education
37 groups among adult men are up to 12 years (Caselli et al. 2014). It has also
38 been shown that the overall decline in life expectancy that Russian men ex-
39 perience over the 1990s was driven by a strong decline among the lower
40 education groups whereas the highest groups continued to enjoy moderate
41 increases (Shkolnikov et al. 2006). In virtually all industrialized countries
42 for which data are available, the education differentials in adult mortality in-
43 creased over time despite improving health care coverage in most countries
44 (Caselli et al. 2014). One explanation for this pattern lies in the increasing
45

2 importance of lifestyle-related factors for which education seems to be more
3 important than the health care system.

4 There is a plausible narrative of causation that is founded in brain re-
5 search. It has been demonstrated that literacy and education in general en-
6 hance the synaptic density in relevant parts of our brains and thus makes
7 us physiologically different for the rest of our lives (Kandel 2007). It has
8 also been shown in a controlled experiment among illiterate Indian young
9 adult men that the sub-sample who was taught how to read and write
10 had lasting structural changes in their brains after six months of learn-
11 ing that are associated with executive functioning and cognitive abilities
12 (Baker, Salinas, and Eslinger 2012; Blair et al. 2005; Brinch and Galloway
13 2012; Skeide et al. 2017). Neurocognitive and neuroimaging studies have
14 shown strong associations between adaptive changes in the brain and learn-
15 ing experiences in the classroom (Lewis et al. 2009; Welberg 2009). These
16 changes have been shown to be associated with ability for abstract thinking,
17 time preference, and the capacity to plan for the future (Cutler and Lleras-
18 Muney 2010; Kenkel 1991; Van der Pol 2011; Heckman, Humphries, and
19 Veramendi 2017). These cognitive changes and resulting changes in behav-
20 ior have relevance for health outcomes. In this context, it should also be
21 stressed that the large number of studies that focus on natural experiments
22 for the education/health link by looking at changes in compulsory school-
23 ing by a year do not seem to be a promising route because they do not
24 refer to a plausible causal pathway. When young people are forced to stay
25 in school against their will for ten rather than nine years, this is unlikely to
26 result in a significant change in their overall brain functioning that would
27 be expected to have direct health effects. But such relevant changes have
28 been shown for major cognitive transitions such as from illiterate to literate
29 or from compulsory to post-secondary education (Cutler and Lleras-Muney
30 2010; Peters et al. 2006; Skeide et al. 2017).

31 Another concern in the context of assessing functional causality is the
32 possibility of reverse causality, which is also implicit in the notion of simul-
33 taneity. This can pose a difficult challenge if there is reason to assume simul-
34 taneous influences going in both directions. Models of Granger causality can
35 be used to sort out the temporal sequence of possible effects that can then be
36 the basis of causal inference based on the principle that the cause needs to
37 precede the effect. But in the case of education such models are unnecessary
38 because the temporal sequence can be identified *a priori*. Schooling tends to
39 happen early in the life course, and it is only the human capital (education
40 stock) at adult ages that is expected to have consequences for health-related
41 behavior. The time lag between when the schooling happens (education
42 flow) and when the resulting stock of human capital influences health can
43 be many decades and, when we study the education/health differentials for
44 people above age 70, even half a century. Hence, once a proper distinction is
45 made between education flow and educational attainment, there is no way

FIGURE 2 Triangular interactions between education flows and stocks, health, and income



in which period conditions at time t (including income, medical inventions, etc.) can influence the schooling that occurred decades before.

With respect to the possibility that the empirical association is caused by a third factor, different pathways of causation need to be distinguished. Figure 2 shows the possible interactions between education, health, and income in the form of a triangle with education and health at the base and income at the top. The cognitive skills associated with adult educational attainment and a person's health status are closely interwoven and both embodied in individuals. A certain minimum level of physical health is necessary for a child to develop mentally and to be able to attend school. Particularly in developing countries, school absenteeism due to poor health of the children themselves or of family members for whom they must care or for whose lost economic output they must compensate is a serious handicap for improving levels of education and, consequently, the building of key cognitive skills and raising learning outcomes. Similarly, children often bring home from school knowledge, attitudes, and access to information relevant to the health and even survival of themselves and their family members.

Better health and longer lives are not only closely linked to education and cognitive capacity but also in turn directly affect economic growth. A growing literature shows that health is both a direct source of human welfare and a driver of income growth (Bloom and Canning 2008). In particular, three mechanisms have been defined: better health leading to higher labor productivity, better childhood health leading to better school

2 attendance and cognitive development, and a longer expected life span lead-
3 ing to more savings and investment. Empirical studies also find that good
4 health has a positive, sizable, and statistically significant effect on aggre-
5 gate output, even controlling for the experience of the work force (Bloom,
6 Canning, and Sevilla 2003). The WHO-sponsored report of the Commission
7 on Macroeconomics and Health (Sachs 2001) also highlighted the impor-
8 tance of basic health for poverty reduction by showing how the burden of
9 diseases in some low-income countries, especially in sub-Saharan Africa,
10 is a barrier to economic growth and therefore should be addressed in any
11 comprehensive development strategy.

12 Finally, the effect of improving educational attainment and quality of
13 education has long been a part of economic growth theory. Past empirical
14 efforts to demonstrate this effect on the basis of aggregate time series have
15 been hampered by the lack of appropriate data on educational attainment
16 (de la Fuente and Doménech 2006). More recent assessments based on full
17 educational attainment distributions by age cohorts demonstrate the consis-
18 tently positive and significant effect of human capital on economic growth
19 (Crespo Cuaresma, Lutz, and Sanderson 2014; Lutz, Crespo Cuaresma, and
20 Sanderson 2008). The same has been shown with respect to quality of ed-
21 ucation for countries where such data exist (Hanushek and Woessmann
22 2008).

24 Updating the global empirical analysis 25 for 1970–2010

26
27 To explore the association between educational attainment, income, and
28 mortality across time and space, we employ a balanced panel of 174 coun-
29 tries (both developed and developing) over the period 1970–2010 in five-
30 year intervals. Following the logic of Preston’s 1975 and 1980 Preston pa-
31 pers, we first present graphical presentations of the bi-variate relationship
32 between GDP per person and life expectancy, followed by multivariate sta-
33 tistical analyses. We also study the pattern with respect to child mortality.
34 While the dichotomous variable of literacy was the only one available to
35 Preston, we use mean years of schooling of the adult population aged 15
36 and older in the case of life expectancy and of women aged 20–39 in the
37 case of child mortality.

39 Data

40
41 Data were obtained mainly from two sources. Country-level indicators of
42 educational attainment for the years 1970–2015 were extracted from the
43 Wittgenstein Center Data Explorer (WIC 2015). Panel data on income and
44 mortality were obtained from the World Development Indicators (World
45 Bank 2017). Merging the two datasets gave us our panel of data with

2 only a few missing data for some countries in earlier years. Following
3 Preston's original design, income is measured as GDP per person (2010 con-
4 stant USD). For the multivariate analysis, we also performed sensitivity runs
5 where PPP (purchasing power parity) per person was used instead of con-
6 stant 2010 USD.

7

8 Descriptive analysis

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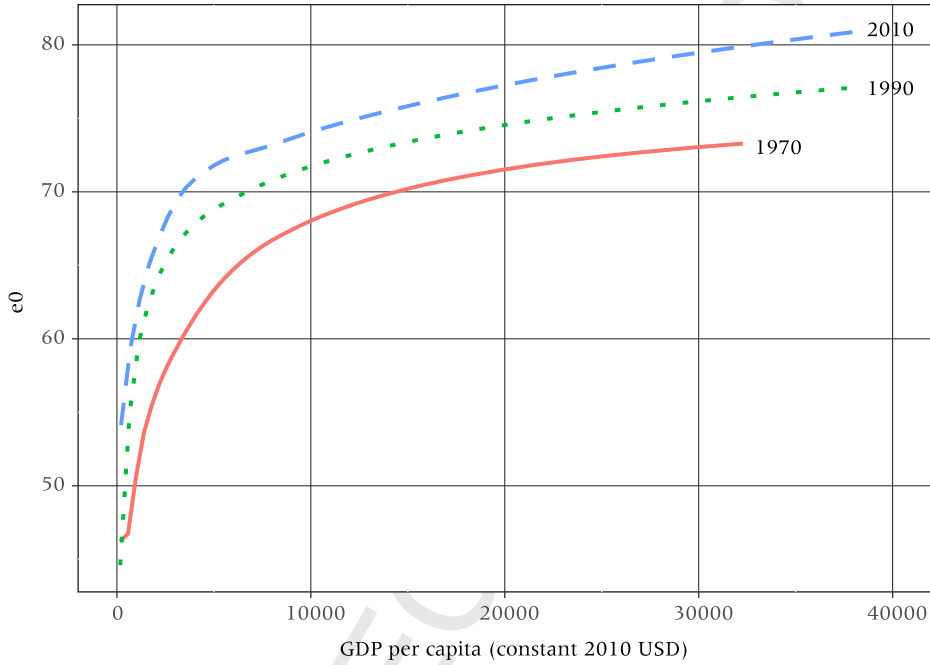
10 Figure 3 presents a visual analysis of the cross-country associations be-
11 tween income, educational attainment, and life expectancy for the years
12 1970, 1990, and 2010, similar to the way in which Preston (1975) plotted
13 life expectancy against GDP per capita for the 1930s and 1960s. In Panel
14 A, the plot of life expectancy at birth against GDP per capita for 1970,
15 1990, and 2010 closely resembles the pattern of the original Preston curve
16 (Figure 1). The curve clearly continues to move upward over time. This,
17 according to Preston, is due to the factors other than income that he as-
18 sumed to be mostly related to medical progress and health care. Panel B
19 of the figure shows an isomorphic curve with the only difference that GDP
20 per person is replaced by mean years of schooling of the adult population
21 (MYS15+). The resulting pattern, however, differs considerably in that ris-
22 ing educational attainment seems to explain rising life expectancy much
23 better than GDP per person. The association looks much more linear with-
24 out a leveling off at higher levels, and there is very little upward shift that
25 would indicate an unexplained gap that needs to be explained by medical
26 progress.

27 Figure 4 shows the corresponding pair of curves in which life ex-
28 pectancy is replaced by child mortality (aged 0–4) in order to see whether
29 the previous pattern also holds for this mortality rate, which is often more
30 easily influenced by targeted health interventions than is overall life ex-
31 pectancy. Here, mean years of schooling of women aged 20–39 is used as
32 the education indicator. In panel A, the plot of child mortality against GDP
33 per person shows essentially the same pattern as for life expectancy vs GDP
34 per person in panel A of Figure 3. Panel B of the figure reveals that the rela-
35 tionship looks much more linear for education than for GDP per person; and
36 for 1970–1990 there is no shift in the curves, with changes in mothers' edu-
37 cation evidently explaining declining child mortality much better than GDP
38 per person. For 1990–2010, however, there is an interesting deviation from
39 this general pattern, as child mortality in the high-mortality/low-education
40 countries declines more rapidly than would be expected from the gains in
41 mothers' education over the same period. Evidently, this is the consequence
42 of the massive international child health interventions in those countries
43 over the past two decades. This shows that, at least with respect to child
44 mortality, concerted public health efforts can lower mortality more than
45 social development alone would predict.

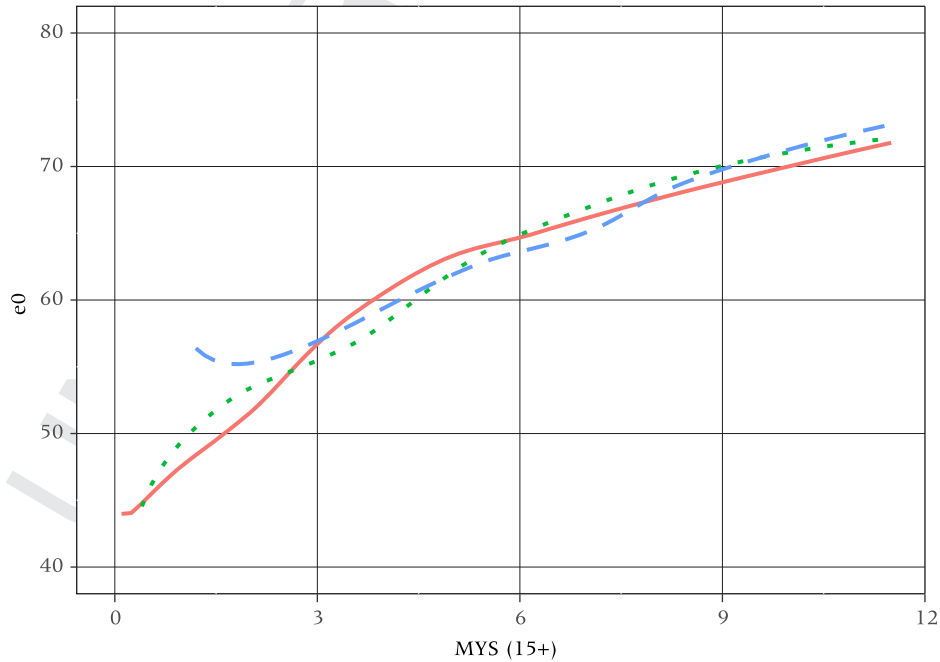
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FIGURE 3 Real GDP per capita (constant USD) and mean years of schooling (MYS) of the adult population aged 15+ plotted against life expectancy at birth in 1970, 1990, and 2010

A. Income vs. e0



B. Education vs. e0

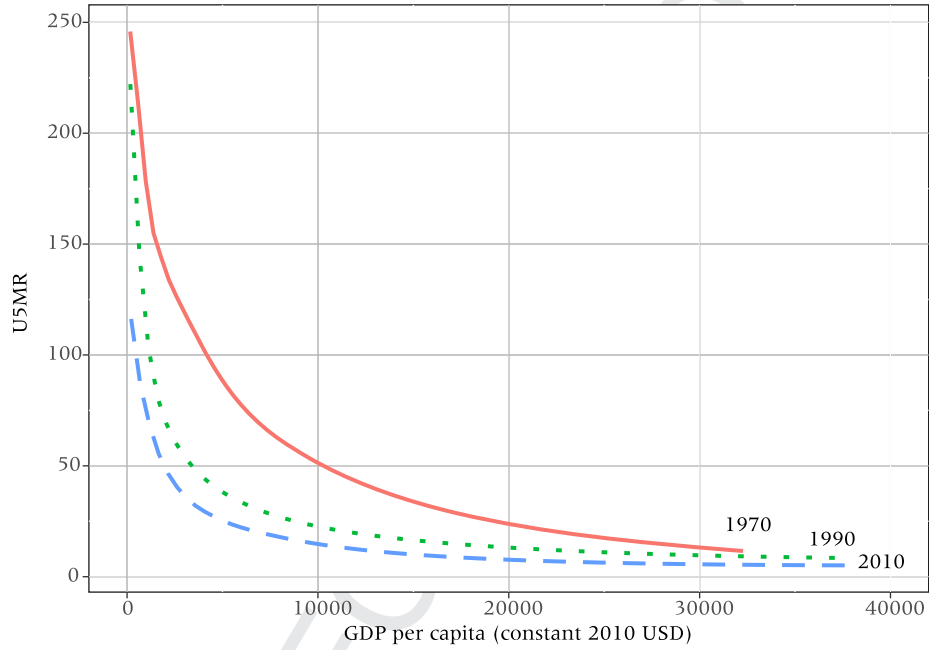


NOTE: 174 countries, lines show fitted splines.

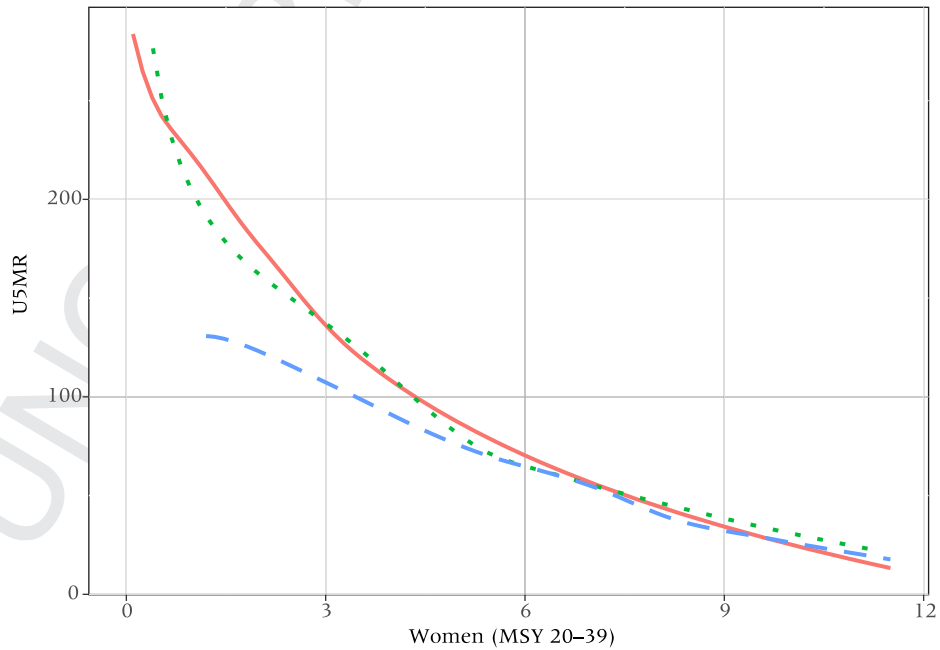
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2 **FIGURE 4 Real GDP per capita (constant USD) and mean years of schooling**
3 **(MYS) of women aged 20–39 plotted against child mortality (U5MR) in 1970,**
4 **1990, and 2010**

5 **A. Income vs. U5MR**



26 **B. Education vs. U5MR**



NOTE: 174 countries, lines show fitted splines.

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2 **TABLE 1 Panel data regressions for the period 1970–2015 (population**
 3 **weighted) with country and time fixed-effects as indicated (standardized**
 4 **coefficients), life expectancy at birth (standardized) as the dependent variable**

5 Variable	(1) Unadjusted	(2) Mutually .Adj.	(3) Main model
6 GDP per person (log)	.914*** (.234)	.307** (.134)	.106 (.091)
8 Education	.994*** (.107)	.855*** (.104)	.395*** (.115)
10 Observations		1276	1,276
11 R-squared		.841	.884
12 Number of countries		149	149
13 Country fixed-effects	Yes	Yes	Yes
14 Year fixed-effects	No	No	Yes

15 ***p<0.01, **p<0.05, *p<0.1.

16 NOTES: GDP per person is in constant 2010 USD and education is the mean years of schooling of the population aged 15+. Robust standard errors in parentheses.

18 Multivariate analysis

19
 20 Multivariate statistical analyses were conducted to explore the relative ef-
 21 fects of education and income in explaining changing global mortality pat-
 22 terns. After conducting a Hausman test to choose between fixed-effect and
 23 random-effect specification for country-level unobserved effects, the pre-
 24 ferred fixed-effect model was fitted and specified as follows:

$$25 M_{it} = U_i + \beta_1 Education_{it} + \beta_2 GDP\ per\ person_{it} + \alpha(t) + \varepsilon_{it} \quad (1)$$

$$26 \alpha(t) = \alpha_1 year_{1970} + \alpha_2 year_{1975} + \alpha_3 year_{1980} + \dots + \alpha_9 year_{2010} + \alpha_{10} \quad (2)$$

27
 28 where M_{it} is the mortality indicator (life expectancy at birth and child mor-
 29 tality) for country i at time t , U_i represent unobservable individual (coun-
 30 try) heterogeneities, $Education_{it}$ is the education indicator (mean years of
 31 schooling for age 15+) for country i at time t , $GDP\ per\ person_{it}$ is GDP per
 32 person (at constant 2010 USD) for country i at time t ; and $\alpha(t)$ controls for
 33 the year fixed-effect and is specified as a binary variable for each year of
 34 observation.
 35

36 A number of different models were estimated and, in addition to the
 37 main models shown in Tables 1 and 2, sensitivity analysis was performed as
 38 described below. For the sake of comparison, all variables were standardized
 39 (Z-scores were calculated) before we fit models. The coefficients are thus
 40 interpreted, within a given country, as the gain in health (in standard de-
 41 viation) for a standard deviation change in the independent variable. First,
 42 we estimated a country fixed-effect panel model by including each explana-
 43 tory variable separately: the unadjusted effect of each predictor is estimated
 44 and shown in column (1). Second, we estimated a model with both edu-
 45 cation and income indicators as shown in column (2). Finally, column (3)

TABLE 2 Panel data regressions for the period 1970–2015 (population weighted) with country and time fixed-effects as indicated (standardized coefficients), under-five child mortality (U5MR) (standardized) as the dependent variable

Variables	(1) Unadjusted	(2) Mutually Adj.	(3) Main model
GDP per person (log)	-.860*** (.029)	-.004 (.118)	.0772 (0.127)
Education	-.923*** (.034)	-1.214*** (.144)	-1.016*** (.130)
Observations		1,257	1,257
R-squared		.824	0.831
Number of countries		147	147
Country fixed-effects	Yes	Yes	Yes
Year fixed-effects	No	no	Yes

***p<0.01, **p<0.05, *p<0.1.

NOTE: GDP per person is in constant 2010 USD and education is the mean years of schooling of women aged 20–39. Robust standard errors in parentheses.

gives the full model with income, education, and both country and period fixed-effects.

The results for life expectancy given in Table 1 show high and significant unadjusted parameters for both income and education after controlling for country fixed-effects, with the standardized education coefficient being somewhat higher. The picture changes substantially when income and education are entered in the same model, with the education coefficient becoming almost three times as large as the one for income. In the full model, which includes time fixed-effects, income becomes insignificant while the education effect remains robust and highly significant.

The results with respect to infant mortality shown in Table 2 are quite similar to those for life expectancy. The difference is that income is already insignificant in the mutually adjusted model in column (2), and the coefficients for women's education (for ages 20–39) in that model and in the full model in column (3) are much higher than the comparable ones in Table 1 for overall mean years of schooling. Sensitivity runs using PPP income instead of constant USD (as in the original Preston paper) only marginally improved the coefficients for income, while the education effects remained robust and highly significant. In addition to the population-weighted regressions shown here, we also ran the model giving every country equal weight. Since our dependent variables here—life expectancy and child mortality—reflect the averages of individual experiences and health-related behaviors, the independent agents whose experience we are studying are individuals rather than countries, which makes population weighting more appropriate. In either case, the results of the unweighted regressions are qualitatively very similar, with the education coefficients slightly lower but still highly significant.

2 These multivariate results strongly confirm what the visual analysis
3 above suggested: raising educational attainment is a much more important
4 driver of increasing life expectancy and falling child mortality than income.
5 As we discussed in the introduction, this finding should have significant im-
6 plications for prioritizing policies aimed at improving health and longevity.
7 Under a global perspective over the last half century, increasing educational
8 attainment clearly has been the key factor in improving health, rather than
9 increasing income as has frequently been claimed.

11 **Summary and conclusions**

13 We revisited the influential 1975 paper by Preston on the relationship be-
14 tween income and life expectancy across most countries of the world for
15 the 1930s and 1960s and extended the analysis to the period 1970–2015.
16 We demonstrated that the distinct pattern identified by Preston, show-
17 ing a strongly concave relationship and an upward shift of the curves,
18 continued over the subsequent half century as assessed at the global
19 level.

20 We then plotted the same kind of relationship replacing GDP per per-
21 son with the mean years of schooling of the adult population to see whether
22 educational attainment could be a better predictor of life expectancy than
23 income. The associations turn out to be very different, with the curves
24 becoming largely linear and overlapping. This suggests that educational
25 attainment is a better predictor in the sense that its effect on life ex-
26 pectancy does not diminish at higher levels and, in particular, it does not
27 leave an unexplained shift over time that has to be explained by other
28 factors.

29 To validate this visual analysis, we conducted multivariate analyses on
30 a balanced panel of 174 countries for 1970–2015, which in addition to GDP
31 per person and mean years of schooling of the adult population included
32 country and period fixed-effects, and we performed sensitivity runs with
33 alternative income indicators and weighting schemes. In all of the models
34 the effect of educational attainment on life expectancy is highly significant
35 in the expected direction, and the standardized coefficients are clearly larger
36 than those of income.

37 To consider the possibility of a different pattern for the determinants
38 of child mortality, we carried out the analysis separately for under-5 mor-
39 tality. Again, for the association with GDP per person there was strong non-
40 linearity and a shift of the curve over time that was particularly pronounced
41 between 1970 and 1990. Viewed in relation to mean years of schooling
42 of women aged 20–39, the relationship again was much more linear with
43 virtually no shift between 1970 and 1990. Between 1990 and 2010 child
44 mortality in the highest-mortality countries declined more rapidly than sug-
45 gested by the gains in mothers' education. This is an indication that massive

efforts by the international community and private donors in recent years to lower child mortality in some of the least developed countries were successful in doing so to a greater extent than would be expected from improving educational attainment alone. This was not equally the case with respect to adult mortality.

Where does this empirical evidence leave us with respect to testing the hypothesis that the empirical association between GDP and life expectancy depicted by the Preston curve and widely assumed in the literature is a spurious one, with education in fact driving both changes? The macro-level evidence presented here strongly supports the view that this is a plausible hypothesis that deserves further elaboration. In our section on causality, we also assessed the different specified criteria for functional causality, including a strong empirical association, a valid narrative of the causal mechanism, and ruling out alternative explanations such as selectivity and reverse causality. This strong aggregate-level finding should be further explored at the individual and community level under different cultural, social, and economic conditions.

Studies such as ours are more easily conducted for child mortality than for adult mortality since the information can be provided in surveys by mothers. Multi-level analyses of the determinants of child mortality across a large number of developing countries have shown that mothers' education at every level of attainment was more important than wealth/income, was the dominating factor at the household and community level, and played a key role at the national level (Pamuk, Fuchs, and Lutz 2011). A comparable individual-level study for adult mortality is much more difficult because there are no consistent data on adult deaths by education, income, and other relevant characteristics. Even in most industrialized countries with efficient vital registration systems, such micro-level analysis will be difficult unless a comprehensive population register exists. With some extra effort, census data that have information on these characteristics can be linked with information about subsequent deaths. Where such matching studies have been conducted, they all show significant mortality differentials by level of education (Caselli et al. 2014). But such studies often lack the income information for comparative analysis. For developing countries, the data challenges are much greater and—possibly except for the cases of demographic surveillance systems—probably insurmountable at present.

The global time series analysis of national data strongly suggests that the apparent positive association between health and income can largely be attributed to increasing educational attainment, which at the same time leads to rising incomes (Lutz, Crespo Cuaresma, and Sanderson 2008) and better health outcomes. While additional individual-level analysis of this issue is needed, the patterns presented here suggest that education should be considered a policy priority for improving global health.

Notes

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