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A FEW METHODOLOGICAL REMARKS

ON OPTIMIZATION RANDOM COST FUNCTIONS

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<u>A Few Methodological Remarks</u> on Optimization Random Cost Functions

Yuri A. Rozanov

Let $f(\alpha, x)$ be a functional of a variable x ϵX , where α is some "unobservable" random parameter with a probability distribution P. Suppose we have to choose some point $x^{O}\epsilon X$, and we like to optimize this procedure in some sense of minimization of $f(\alpha, x)$, x ϵX , with unknown parameter α .

For example, $f(\alpha, x)$ may be a cost function of some economic model concerning future time, say

$$f(\alpha, x) = \sum_{1}^{n} \alpha_{j} x_{j}, \quad x = (x_{1}, \dots, x_{n}) \varepsilon X, \quad (1)$$

where X is a given convex set in n-dimensional vector space formed with inequalities

$$\sum_{i=1}^{n} \alpha_{i} x_{j} \geq b_{i} , \quad i = 1, ..., m$$
 (2)

(including $x_j \ge 0$; j = 1, ..., n), and $\alpha = (\alpha_1, ..., \alpha_n)$ is a vector of "cost coefficients," which are expected to take 'values with some probability distribution $P(\cdot | \delta)$ under conditions of some given data δ .

Sometimes one uses a criterion based on minimization of mean value $Ef(\alpha, x)$, x ϵX , and considers x^{O} as the optimal

point if

$$Ef(\alpha, x^{O}) = \min_{x \in X} Ef(\alpha, x) .$$
(3)

This criterion looks quite reasonable if one is going to deal with a <u>big number N of similar models</u>, and the total cost function can be approximately described (according to central limit theorem) as

$$\sum_{k=1}^{N} f(\alpha_{k}, x) \approx [Ef(\alpha, x)] \cdot N + \Theta \sqrt{N} ,$$

where Θ is a random (normal) variable with mean zero and variance $\sigma^2(\mathbf{x}) = Df(\alpha, \mathbf{x})$. But if you have to put in a big investment only once, then mean value criterion may not work well; moreover, the minimum point \mathbf{x}^{O} of mean value function Ef(α, \mathbf{x}), $\mathbf{x} \in \mathbf{X}$, can be the maximum point of the cost function f(α, \mathbf{x}), $\mathbf{x} \in \mathbf{X}$, with a great probability.

In order to make this obvious remark clearer, let us mention a model of a non-symmetric coin game with two outcomes: $\alpha = \alpha_1, \alpha_2$, which takes place with corresponding probabilites $p_1, p_2 = 1 - p_1$, and cost function is $f(\alpha, x)$ with $x = x_1, x_2$. One has to pay $f_{ij} = f(\alpha_i, x_j)$ under the outcome α_i if he chooses in advance the strategy x_j (i, j = 1, 2). Suppose $f_{ij} = C$ (i \neq j), where C is the all gambler capital (so he will lose this capital C under the strategy x_j if it be the outcome α_i , $i \neq j$), and $f_{ii} = -M_iC$ (he will increase the initial capital C in M, times). The mean value function is

$$Ef(\alpha, x) = \begin{cases} C(-M_1p_1 + p_2) & \text{if } x = x_1 \\ \\ C(p_1-M_2p_2) & \text{if } x = x_2 \end{cases}$$

Suppose the outcome α_1 takes place with a great probability p_1 (say $p_1 = 0.999$) and M_2 is so big that

$$p_1 - M_2 p_2 < -M_1 p_1 + p_2$$

Using mean value criterion, we obtain $x^{\circ} = x_2$ as the optimal point, but obviously this is a very foolish strategy, except in the case when one should very much like to lose his capital (because it will be with the great probability 0.999). Another similar example: suppose the cost function is

$$f(\alpha, \mathbf{x}) = \begin{cases} \alpha_{10} + \alpha_{11} \mathbf{x} & \text{with probability } \mathbf{p}_1 \\ \\ \alpha_{20} + \alpha_{21} \mathbf{x} & \text{with probability } \mathbf{p}_2 = 1 - \mathbf{p}_1 \end{cases}$$

(say $p_1 = 0.999$, $p_2 = 0.001$) where $0 \le x \le 1$ and the cost coefficients α_{11}, α_{21} are such that $\alpha_{11} > 0$; $\alpha_{11}p_1 + \alpha_{21}p_2 < 0$.

Using mean value criterion, we have to choose $x^{\circ} = 1$, though with the great probability p_1 ($p_1 = 0.999$) it will be the <u>maximum</u> point (see Fig. 1) of the actual cost function $f(\alpha, x)$, $0 \le x \le 1$.

Concerning the mean value type criterion, we wish to say some other things. It is very easy to realize that one may prefer a random variable $n_1 = f(\alpha, x_1)$ in comparison to



FIGURE 1

$$F_1(y) = P\{n_1 \le y\} \ge \{n_2 \le y\} = F_2(y)$$
.

Of course, there may be a few, in some sense, crucial points $y = y_1, \ldots, y_n$. Suppose it is possible to estimate "an importance" of these points with the corresponding values u(y), $y = y_1, \ldots, y_n$ in such a way that one prefers η_1 (as compared to η_2) if

$$\sum_{k} F_{1}(y_{k}) u(y_{k}) \geq \sum_{k} F_{2}(y_{k}) u(y_{k})$$

The preference relation can be rewritten in the form

$$\int F_1(y) \ dU(y) \geq \int F_2(y) \ dU(y) ,$$

where

$$U(y) = \sum_{\substack{k \leq y \\ y_k \leq y}} u(y_k), \quad -\infty < y < \infty.$$

Because for any distribution function $F(y)(F(-\infty) = 0, F(\infty) = 1)$ we have

$$\int F(y) dU(y) = - \int U(y) dF(y) + U(\infty) ,$$

the preference criterion can be represented in the form

$$EU(\xi_1) \leq EU(\xi_2) , \qquad (4)$$

where E(•) is the corresponding mean value.

One can consider (4) for arbitrary distribution type function U(y), $-\infty < y < \infty$ as the general <u>mean value criterion</u>. Obviously, if the corresponding density u(y), $-\infty < y < \infty$ is <u>positive</u>, then U(y), $-\infty < y < \infty$ is a monotone increasing function. Besides, if for any $y_1 \le y_2$ on some interval we consider y_1 as "more important" in comparison with y_2 , more precisely if

$$u(y_1) \ge u(y_2)$$
 , $y_1 \le y_2$,

i.e. the density u(y), $x \in I$ is a monotone decreasing function on the interval I, then the preference function U(y), $y \in I$, is convex (see Fig. 2).

We are going to suggest below a few other types of criteria of optimization for random cost functions.

1. Let $f(\alpha, x)$, $x \in X$ be a cost function which depends on a random parameter α . Suppose for some <u>acceptable</u> cost value C we can neglect a probability that the actual cost will exceed C. Suppose that minimal (random) cost

$$C(\alpha) = \min f(\alpha, x)$$

xeX

has a probability distribution with a rather <u>small range</u> and corresponding minimum point $\xi \in X$:

$$f(\alpha,\xi) = \min_{x \in X} f(\alpha,x)$$

has a discrete distribution (maybe with a very big dispersion).



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FIGURE 2

It seems quite reasonable to take a risk to choose such point $x^{O} \epsilon X$ for which

$$P\{f(\alpha, x^{O}) = C(\alpha)\} = \max_{x \in X} P\{f(\alpha, x) = C(\alpha)\} .$$
(5)

Note that if the probability in the relation (5) equals to 1, in other words, there is a point $x^{O} \epsilon X$ for which

$$f(\alpha, x^{O}) = \min f(\alpha, x)$$
 with probability 1 , $x \in X$

then our criterion gives the usual minimum of cost function. Let us consider the linear cost function

$$\mathbf{f}(\alpha,\mathbf{x}) = \sum_{j=1}^{n} \alpha_{j} \mathbf{x}_{j}$$

of $x = (x_1, \ldots, x_n) \in X$, where $\alpha = (\alpha_1, \ldots, \alpha_n)$ is the random vector with a given probability distribution P, and X is a simplex in n-dimensional vector space of the type (2):

$$\sum_{i=1}^{n} a_{ij} x_{j} \ge b_{i} ; \quad i = 1, \dots, m$$

Denote x^1, \ldots, x^N extreme points of simplex X. As well known, a <u>minimum</u> point $\xi \in X$ (ξ depends on α) can be chosen among x^1, \ldots, x^N , so $x^\circ = x^1, \ldots, x^N$ is the optimal point in the sense of the criterion (5) if

$$P\{\xi = \mathbf{x}^{\mathsf{O}}\} = \max_{\substack{1 \le k \le \mathsf{N}}} P\{\xi = \mathbf{x}^{\mathsf{k}}\} .$$
(6)

Thus, the problem is to find all probabilities*

$$P_k = P\{\xi = x^k\}$$
; $k = 1,...,N$

and to choose the optimal x^{O} as the point among x^{k} ; k = 1, ..., N, with the greatest probability P_{k} ; k = 1, ..., N.

We have $P_k = P(Y^k)$ where Y^k is the set of all vectors $y = (y_1, \dots, y_n)$ for which the corresponding linear function

$$f(y,x) = \sum_{j=1}^{n} y_j x_j$$
, $x \in X$

has x^k as the minimum point:

$$f(y,x^k) = \min_{x \in X} f(y,x) .$$

In order to make our elementary consideration more clear, let us shift x^k to the origin point x = 0. Obviously, the extreme point $x^k = 0$ gives a minimum of f(y,x), $x \in X$, iff

$$\sum_{j=1}^{n} y_{j} x_{j} \geq 0 \text{ for all } x \in X$$

(in other words, iff the vector $y = (y_1, \dots, y_n)$ belongs to socalled <u>polar cone</u>).

Let us take all hyperplains

$$\sum_{1}^{n} a_{ij} x_{j} = b_{i} , i \varepsilon I_{k}$$
(7)

* Note the events $\{\xi = x^k\}$; k = 1, ..., N generally are not disjoined and $\sum_{k=1}^{N} P_k$ not necessary equals to 1.

--see (2)--containing the extrème point x^k . (In the case $x^k = 0$ we have $b_i = 0$, $i \in I_k$.) Let us introduce a cone

$$x^{k} = \bigcap_{i \in I_{k}} \{x: \sum_{i=1}^{n} a_{ij} x_{j} \ge 0\} .$$

The corresponding polar cone is exactly the set Y_o^k of all vectors $y = (y_1, \dots, y_n)$ such that $\sum_{j=1}^n y_j x_j \ge 0$, $x \in X^k$ (see Fig. 3). This polar cone Y_o^k is formed by all linear combinations

$$y = \sum_{i \in I_{k}} \lambda_{i} a_{i} ; \quad \lambda_{i} \ge 0$$
(8)

of the vectors $a_i = (a_{i1}, \dots, a_{in})$, $i \in I_k$ because a dual polar cone for the set of all vectors (8) coincides with x^k : obviously,

$$\sum y_{j}x_{j} = \sum_{i \in I_{k}} \lambda_{i} (\sum a_{ij}x_{j}) \ge 0$$

for all $\lambda_{i} \geq 0$, iff $x \in x^{k}$. (See, for example, duality theorem in [1].) Thus, $\underline{y^{k} = x^{k} + \underline{y}_{0}^{k}}$ is the set of all vectors

$$y = x^{k} + \sum_{i \in I_{k}} \lambda_{i} a_{i} , \quad \lambda_{i} \geq 0 , \qquad (9)$$

where $a_i = (a_{i1}, \dots, a_{in})$ are all vectors such that for $x = x^k$ at the relations (2) we have strict equalities, and the optimal point can be found among x^k , $k = 1, \dots, N$ as a point with maximum probability

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$$P(Y_k) = P\{\alpha \in Y_k\}$$
; $k = 1,...,N$. (10)

2. Suppose, as above, there is the <u>acceptable</u> cost, which can be exceeded only with a corresponding small probability, but the situation is different in the sense that the range of the minimum cost distribution is considerably big. (For example, the minimum point $\xi = x^1, x^2$ can be distributed with almost equal probabilities $P_1 > P_2$, but corresponding cost values are such that $f(\alpha, x^1) >> f(\alpha, x^2)$, so there is no reason to choose the point x^1 with the greatest probability P_1 as optimum.)

Suppose that one is going to risk in order to make the cost value less than some level C_0 . (Probability P { $C(\alpha) \leq C_0$ } has to be considerably big.) Then one can choose <u>optimal</u> <u>point</u> $x^0 \in X$ in the sense that

$$P \{f(\alpha, x_n^{O}) \leq C_{O}\} = \max_{x \in X} P \{f(\alpha, x) \leq C_{O}\} .$$
(11)

This criterion is of mean value type (4) concerning a new cost function $EU(f(\alpha, x))$, xeX where

$$U(y) = \begin{cases} 1 & \text{if } y \leq C_{o} \\ 0 & \text{if } y > C_{o} \end{cases}$$

namely,

$$EU(f(\alpha, x^{O})) = \min EU(f(\alpha, X)) .$$
(12)
x \varepsilon X

(Note it is impossible to restrict "y" in order to deal with the convex function U(y), yeI.)



FIGURE 4

3. Suppose, now, there is a good deal of risk to pay a big amount if we use "extreme strategy" x^{O} of types (5) or (11), because with considerably big probability, cost value $f(\alpha, x^{O})$ may be too much. Suppose one should like to prevent a danger of dealing with the "almost worst" outcome α , and the problem is to find optimal strategy against "very clever random enemy." In this situation, the following criterion seems quite reasonable (similar to the <u>minimax principal</u> of game theory).

Namely, suppose one agrees (roughly speaking) to risk only with a small probability $\varepsilon \ge 0$. Let C(x) be the " ε -quantil" for the random variable f(α, x):

$$C(\mathbf{x}) = \min C | P\{f(\alpha, \mathbf{x}) \leq C\} \leq 1 - \varepsilon .$$
 (13)

$$C(x^{O}) = \min_{x \in X} C(x) .$$
(14)

In the case of $\varepsilon = 0$, our criterion of optimality coincides with well known <u>minimax principal</u> of the game theory, which was mentioned above, because if $\varepsilon = 0$, then

$$C(x) = \sup_{\alpha} f(\alpha, x)^{\circ}$$
.

(We mean so-called essential sup $f(\alpha, x)$ concerning the probability distribution P of the random variable α .)

For the linear cost function (1) with the coefficients $\alpha = (\alpha_1, \dots, \alpha_n)$ which are weakly dependent, one can expect the random variable $f(\alpha, x) = \sum_{j=1}^{n} \alpha_j x_j$ is <u>normally distributed</u> (due to the central limit theorem) with a mean value

$$(\mathbf{c},\mathbf{x}) = \sum_{1}^{n} \mathbf{c}_{j}\mathbf{x}_{j}$$

and variance

$$|| \sigma^{\frac{1}{2}} \mathbf{x} ||^{2} = \sum_{i=1}^{n} \sigma_{ij} \mathbf{x}_{i} \mathbf{x}_{j}$$

$$(c_i = E\alpha_i; \sigma_{ij} = E(\alpha_i - c_j)(\alpha_j - c_j); i, j = 1, \dots, n).$$

If it holds true, then

$$C(\mathbf{x}) = \sum_{i=1}^{n} c_{i} \mathbf{x}_{i} + \mathbf{y}_{\varepsilon} \left(\sum_{i=1}^{n} \sigma_{ij} \mathbf{x}_{i} \mathbf{x}_{j}\right)^{\frac{1}{2}}, \quad \mathbf{x} \in \mathbb{X} ,$$

where $\boldsymbol{\gamma}_{\epsilon}$ denotes $\epsilon\text{-quantil for the standard normal distribution:$

$$\sqrt{2\pi} \int_{Y_{\epsilon}}^{\infty} e^{-y^{2}/2} dy = \epsilon$$
.

This function

•

$$C(x) = (c,x) + Y_{\varepsilon} || \sigma^{\frac{1}{2}} x || , x \varepsilon X$$

(where $\sigma^{\frac{1}{2}}$ means the square root of the positive matrix $\{\sigma_{ij}\}$) for $y_{\epsilon} > 0$ is <u>concave</u> because

$$\| \sigma^{\frac{1}{2}} \frac{x_{1} + x_{2}}{2} \| \leq \frac{1}{2} \left(\| \sigma^{\frac{1}{2}} x_{1} \| + \| \sigma^{\frac{1}{2}} x_{2} \| \right)$$

and the minimum point x^{O} can be found with well known concave programming methods. (See, for example, [1].)

References

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